

Frequently Asked Questions

1. Define exponential distribution.

Answer:

A continuous random variable X is said to have an exponential distribution with parameter θ if its pdf is given by, $f(x) = \theta e^{-x/\theta}$, $x > 0$.

And we write $X \sim \text{exp}(\theta)$

2. Define exponential distribution with mean θ

Answer:

A continuous random variable X is said to have an exponential distribution

with mean θ if its pdf is given by, $f(x) = \frac{1}{\theta} e^{-x/\theta}$, $x > 0$

3. Obtain mgf of exponential distribution.

Answer:

If $X \sim \text{exp}(\theta)$, then mgf is given by,

$$\begin{aligned} M_X(t) &= E(e^{tx}) = \int_0^{\infty} e^{tx} f(x) dx = \theta \int_0^{\infty} e^{tx} e^{-x/\theta} dx \\ &= \theta \int_0^{\infty} e^{-(\theta-t)x} dx = \theta \left[\frac{e^{-x(\theta-t)}}{-(\theta-t)} \right]_0^{\infty} = \frac{\theta}{\theta-t} = \frac{1}{1-t/\theta} = (1-t/\theta)^{-1} \end{aligned}$$

provided $|t/\theta| < 1$

4. Derive mean of the distribution from mgf using differentiation.

Answer:

We know that $M_X(t) = [1 - (t/\theta)]^{-1}$

$$E(X) = \frac{d}{dt} M_X(t) \Big|_{t=0} = \frac{d}{dt} \left(1 - \frac{t}{\theta} \right)^{-1} \Big|_{t=0} = - \left(1 - \frac{t}{\theta} \right)^{-2} \times \left(-\frac{1}{\theta} \right) \Big|_{t=0} = \frac{1}{\theta}$$

5. Obtain variance of the exponential distribution from mgf using the method of differentiation.

Answer:

Variance of the distribution is given by, $V(X) = E(X^2) - [E(X)]^2$

$$E(X) = \frac{d}{dt} M_X(t) \Big|_{t=0} = \frac{d}{dt} \left(1 - \frac{t}{\theta} \right)^{-1} \Big|_{t=0} = - \left(1 - \frac{t}{\theta} \right)^{-2} \times \left(-\frac{1}{\theta} \right) \Big|_{t=0} = \frac{1}{\theta}$$

$$\begin{aligned} E(X^2) &= \frac{d^2}{dt^2} M_X(t) \Big|_{t=0} = \frac{1}{\theta} \frac{d}{dt} \left(1 - \frac{t}{\theta} \right)^{-2} \Big|_{t=0} \\ &= \left(\frac{1}{\theta} \right) \left\{ -2 \times \left(1 - \frac{t}{\theta} \right)^{-3} \times \left(-\frac{1}{\theta} \right) \Big|_{t=0} \right\} = \frac{2}{\theta^2} \end{aligned}$$

$$V(X) = \frac{2}{\theta^2} - \left(\frac{1}{\theta} \right)^2 = \frac{1}{\theta^2}$$

6. Write the relationship between mean and variance of exponential distribution.

Answer:

If $X \sim \text{exp}(\theta)$, then mean = $1/\theta$ and variance = $1/\theta^2$

I.e. Variance = $(1/\theta)(1/\theta) = \text{mean}/\theta$

Therefore Variance > Mean, if $0 < \theta < 1$

Variance = Mean if $\theta = 1$ and

Variance < Mean if $\theta > 1$

Hence for the exponential distribution,

Variance >, = or < Mean for different values of the parameter.

7. Obtain the expression for mean and variance by equating the coefficient of $t^r/r!$ in mgf.

Answer:

Let us find mean and variance by equating the coefficient of $t^r/r!$. Consider the mgf,

$$M_X(t) = \left(1 - \frac{t}{\theta}\right)^{-1} = 1 + \frac{t}{\theta} + \left(\frac{t}{\theta}\right)^2 + \dots$$

Now $E(X^r) = \text{coefficient of } t^r/r!$.

Therefore $E(X) = \text{coefficient of } t/1! = 1/\theta$

$E(X^2) = \text{Coefficient of } t^2/2! = 2/\theta^2$

Hence $V(X) = E(X^2) - [E(X)]^2$

$$= 2/\theta^2 - (1/\theta)^2$$

$$= 1/\theta^2$$

8. Find Cumulant generating function.

Answer:

We know that Cumulant generating function is given by

$$K_X(t) = \log(M_X(t)) = \log[1 - (t/\theta)]^{-1}$$

$$= -1 \log[1 - (t/\theta)]$$

$$= -1 \left[\frac{t}{\theta} - \frac{\left(\frac{t}{\theta}\right)^2}{2} - \frac{\left(\frac{t}{\theta}\right)^3}{3} - \frac{\left(\frac{t}{\theta}\right)^4}{4} - \dots \right]$$

$$= \frac{t}{\theta} + \frac{\left(\frac{t}{\theta}\right)^2}{2} + \frac{\left(\frac{t}{\theta}\right)^3}{3} + \frac{\left(\frac{t}{\theta}\right)^4}{4} + \dots$$

9. Obtain the first four Cumulants from cgf.

Answer:

$K_1 = \text{coefficient of } t/1! = 1/\theta$

$K_2 = \text{coefficient of } t^2/2! = 1/\theta^2$

$K_3 = \text{coefficient of } t^3/3! = 2/\theta^3 = \mu_3$

$K_4 = \text{coefficient of } t^4/4! = 6/\theta^4$

10. Comment on the nature of an exponential distribution.

Answer:

Let us find coefficient of skewness

$$\gamma_1 = \frac{\mu_3}{\mu_2^{3/2}} = \frac{2/\theta^3}{(1/\theta^2)^{3/2}} = 2 > 0$$

Coefficient of kurtosis is given by,

$$\beta_2 = \frac{\mu_4}{\mu_2^2} = \frac{9/\theta^4}{(1/\theta^2)^2} = 9 > 3$$

Hence exponential distribution is positively skewed and has leptokurtic curve.

11. Find median of an exponential distribution.

Answer:

Median divided the distribution into two equal parts. Hence if M is the median

of the distribution then $\int_0^M f(x)dx = \int_M^\infty f(x)dx = \frac{1}{2}$

We can consider any one of the integral. Let us consider the first.

$$\theta \int_0^M e^{-x\theta} dx = \frac{1}{2}$$

$$\Rightarrow - \left| e^{-x\theta} \right|_0^M = \frac{1}{2}$$

$$\Rightarrow e^{-M\theta} = \frac{1}{2}$$

$$\Rightarrow M = \log_e(2) / \theta$$

12. What is the mode of an exponential distribution?

Answer:

Observe that exponential distribution is strictly a decreasing function, $f(x)$ is maximum at $x=0$. Hence mode of the distribution is zero.

13. State and prove memoryless property of an exponential distribution.

Answer:

Exponential distribution 'lacks memory', i.e. if X has an exponential distribution, then for every constant $a \geq 0$, one has

$P(Y \leq x / X \geq a) = P(X \leq x)$ for all x , where $Y = X - a$

Proof:

We have $P(Y \leq x \cap X \geq a) = P(X - a \leq x \cap X \geq a)$

$$= P(X \leq x + a \cap X \geq a)$$

$$= P(a \leq X \leq a + x)$$

$$= \theta \int_a^{a+x} e^{-\theta x} dx = e^{-a\theta} (1 - e^{-\theta x})$$

$$\text{And } P(X \geq a) = \theta \int_a^\infty e^{-\theta x} dx = e^{-a\theta}$$

$$P(Y \leq x / X \geq a) = \frac{P(Y \leq x \cap X \geq a)}{P(X \geq a)} = \frac{e^{-a\theta} (1 - e^{-\theta x})}{e^{-a\theta}} = (1 - e^{-\theta x})$$

$$\text{Also } P(X \leq x) = \theta \int_0^x e^{-\theta x} dx = 1 - e^{-\theta x}$$

Hence,

$$P(Y \leq x / X \geq a) = P(X \leq x)$$

i.e. exponential distribution lacks memory.

14. Show that If $X_i, i = 1, 2, n$, are independent exponential r.v.s with parameter θ_i . Let $Z = \min(X_1 \dots X_n)$, then $Z \sim \exp(\sum \theta_i)$.

Answer:

Since X_i 's are independent exponential r.v.s with parameter θ_i , pdf is given by $f(x_i) = \theta_i e^{-x\theta_i}, x_i > 0$

To prove above result first we consider

$$P(Z > x) = P[\min(X_1 \dots X_n) > x] \\ = P[X_1 > x, X_2 > x \dots X_n > x]$$

Since the random variables are independent,

$$P(Z > x) = P[X_1 > x]P[X_2 > x] \dots P[X_n > x] \\ = e^{-\theta_1 x} e^{-\theta_2 x} \dots e^{-\theta_n x} = \prod_{i=1}^n e^{-\theta_i x} = e^{-(\sum_{i=1}^n \theta_i)x}, \text{ which is the mgf of exponential}$$

distribution with parameter $\sum \theta_i$. Hence by uniqueness theorem of mgf, $Z \sim \exp(\sum \theta_i)$

15. Suppose that the amount of time one spends in a bank is exponentially distributed with mean 10 minutes, what is the probability that a customer will spend more than 15 minutes in the bank? What is the probability that a customer will spend more than 15 minutes in the bank given that he is still in the bank after 10 minutes?

Answer:

Let X denote amount of time one spends in a bank. Hence $X \sim \exp(\theta)$

Mean is 10 minutes. Hence $\theta = 1/10 = 0.1$

Therefore pdf is given by, $f(x) = \theta e^{-x\theta}, x > 0$.

By substituting θ , we get $f(x) = (0.1)e^{-0.1x}, x > 0$

Consider

$P(\text{Customer will spend more than 15 minutes})$

$$= P(X > 15) = 0.1 \int_{15}^{\infty} e^{-0.1x} dx = e^{-0.1 \times 15} = 0.22$$

Now we need to find the probability that customer will spend more than 15 minutes in the bank given he is still in the bank after 10 minutes

i.e., $P(X > 15 | X > 10)$

Using memory less property we can write,

$$P(X > 15 | X > 10) = P(X > 5) = e^{-0.1 \times 5} = 0.604$$